Multisensor data fusion for OD matrix estimation

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Abstract Knowledge of traf@ demand at a junction is crucial for most transport systems. Generally, it is represented by an origin-destination (OD) matrix, where each element is a volume of vehicle œow between one of the OD pair of zones of a junction. This paper introduces a new method for a short-time estimation of OD matrices at a signalised junction with a complex structure and Officed out with video cameras. The estimation is made by taking into account the trafŒ lights and by fusion of different multisensor trafted data, which are subject to imprecision and uncertainty. The method proceeds in two steps. First, vehicle conservation law, expressed in terms of undertermined system of equations, is built dynamically for a short-time period. Second, four approaches, that overcome the system indetermination and model the data imperfection, are proposed to provide the best and unique estimates of the OD matrix. An experimental study has been made with data collected at the real signalised junction.

Index Terms Fuzzy sets, fuzzy modelling, fuzzy linear programming, origin-destination matrix estimation.

I. INTRODUCTION

Information on trafted demand in a road network can be represented using the origin-destination (OD) matrix \mathbf{B} . For a junction, each element b_{ij} of such a matrix is a proportion of the ∞ of vehicles that come from entrance (origin) i and go to exit (destination) j. The matrix elements are also called the OD ∞ w rates.

OD matrix estimated for a junction is a fundamental component for most of transport systems. Knowledge about trafted demand at a junction is generally used for trafted planning purposes, such as building of new links, rearrangement of turning lanes, change of the capacity of income/outcome links and of the speed limits, construction of roundabouts. Another important application of the junction OD matrix is to improve estimation of trafted demand at a whole road network. Thus, it is used in such trafted models as the model of vehicle s route choices designed for solving the trafted assignment problem.

Whenever a congestion occurs because of accident, meteorological disturbance or road maintenance work, it is important to assume fast and efteient measures to eliminate it. Therefore, information on destinations of travellers at a junction helps to develop and improve the trafte control strategies. The OD matrix is also employed for traveller information systems and validation of different trafte models.

Because of temporal variability of the trafte demand, most of these systems require real-time knowledge about OD matrix.

Therefore, the estimation period of the OD ∞ was rates has to be as short as possible, like a trafæ light cycle (duration of green-amber-red sequence) at a signalised junction. Such a short-time estimated OD matrix is a key element of the diagnostic system for signalised junctions, developed at INRETS/GRETIA [1]. This system compares the impacts of different trafæ control strategies formulated, for instance, in terms of CO₂ and pollutant emissions.

Junction OD matrix is deduced from vehicle counts made on each income and outcome link of the junction during a given time interval. These counts are usually provided by magnetic loops embedded in the road surfaces and sensitive to metallic masses. Since the loops are installed quite far from the entrances and exits of a junction, the OD matrix estimation period should be large enough in order to take into account the vehicle travel time between an OD pair of sensors and thus to respect the vehicle conservation law. In the case of a signalised junction, the travel time across the junction and the input and output œow volumes during a trafte cycle remain unknown. Unlike the electromagnetic loops, video sensors are able to provide the vehicle counts at any point of the junction, in particular, at the beginning of each entrance and each exit of the junction, which make possible to estimate an OD matrix at a short period of time.

The previous research works on OD matrix estimation at a junction generally yield estimates of the OD αων rates at a large-time period. Some of these methods are based on the information minimisation principle [2], on maximisation of likelihood [3] or on Bayesian inference [4]. Other methods present the estimation problem as a constraint optimisation problem [5], [6], [7]. Only a few methods can provide a short-time estimated OD αων rates, such as a recursive estimation method [8], [9] and methods based on the use of traft lights [10], [11], [12]. However, they can only be applied for simple two-road-crossing junctions and, respectively, for junctions with two-phase traft control strategy. Moreover, none of the cited methods considers the redundancy, uncertainty and imprecision of the data.

This paper focuses on the short-time estimation of OD matrices at a signalised junction with a complex structure and Œted out with video cameras. In particular, we are interested in the use of trafŒ lights and imprecise and uncertain trafŒ measurements issued from video cameras. These measure-

ments, provided every second, are the vehicle counts made on each entrance and exit of the junction and the number of vehicles stopped at each inner section of the junction. The main idea of the proposed method consists in constructing at each trafŒ light cycle a set of (fuzzy) relationships between the sequences of exit and entrance œw counts. A model of the vehicle œws crossing the junction is built via High-Level Petri Nets (HLPN) [13]. This model indicates every second a set of vehicle œws which could be present at the junction, the beginning and the end of these œws. It makes possible to draw dynamically a conservation law of vehicles, represented by an underdetermined system of equations whose unknowns are the OD œw rates.

Four different approaches, that overcome the inconsistency of the system and model data imperfections, are proposed to estimate the OD ∞ rates: ordinary least squares and linear programming (LP) methods based on the use of historical knowledge about the OD ∞ rates, LP and fuzzy LP methods based on a modelling of the imprecision of the data using, respectively, intervals or fuzzy numbers. An experimental study of the proposed methods has been made with the real data collected in an experimental site equipped with video cameras and trafæ light controller. These tests demonstrated that the presented dynamic method yields accurate results.

The rest of the paper is organised as follows. Section 2 describes the real data. Section 3 brieze introduces the model of OD zows crossing the junction and states the problem. Section 4 and 5 proposes four methods for OD matrix estimation and shows the results. The last section concludes the paper and proposes the further lines of research.

II. DATA DESCRIPTION

The experimental site is an isolated signalised junction of two double-lane roads situated in the south suburb of Paris (cf Fig. 1). The main road B-D which connects the suburbs to Paris has a high traft volume, whereas the road A-C has a lower traft volume. Traft lights control four incoming links and four inner zones. Note that right-turning vehicles use special lanes and are not taken into account in this study. Only eight OD α was are statistically signiteant: (AC, AD), (BD, BB), (CB), (DB, DC, DD).

Eight video cameras are installed at the junction in order to capture all the entrance and exit links and the inner zones. The location, height and angle of each camera depend on the geometry of the junction and are chosen to favour the measurement of space traft parameters such as queue length on incoming links. The camera views are analysed in real time using image processing techniques developed at INRETS [14]. They provide several measurements every second:

- $\mathcal{X}_i(\tau)$ vehicle counts measured at the end of an entrance i at second τ ($i = 1, \ldots, n$),
- · y_j(τ) number of vehicles that have passed through the beginning of the exit j at second τ (j = 1,...,m),
- $\mathcal{Z}_k(\tau)$ number of stopped vehicles at inner zone k at second τ $(k=1,\ldots,p)$,

where n, m and p are the numbers of entrances, exits and inner zones respectively. Here, a $traf CE \ light \ cycle$ is a period of time between two sequential onsets of the red light on the entrance Bf.

The temporal series can be represented graphically. Figure 2 presents an example of measurements of the vehicle cow arrived from the entrance Af and crossing some successive zones toward exits Cf and Df. Here vertical lines (solid and stippled) indicate the onsets of the red and green light respectively. The grey squares point out the vehicles counts at the corresponding entrance and exits. The white squares represent the number of vehicles stopped at inner zones. It can be seen that the volumes of incoming and exit cows are equal, that is to say 4/5 of vehicles arrived from the entrance Af have leaved the junction by exit Cf and 1/5 by exit Df. Actually, one vehicle from OD cow ADf has stopped twice in the inner zones 1f and 2f because of the red light in zone 2f. The Figure 2 demonstrates that the measurements are captured in inner zones with a delay with regard to the actual trafte situation.

It is worth noting that generally the vehicle counts are imprecise and erroneous so that the vehicle conservation law is not respected. The vehicles counts are more imprecise than the measure of occupation rates of inner zones by stopped vehicles.

The low quality of the measurement of trafte parameters is due to many phenomena. Trafte conditions (peak or off-peak periods) are the reason for many trafte count errors. If the trafte cow is heavy, the gaps between vehicles are small and it is difteult to distinguish these gaps on the video images. Thus the number of vehicles measured is lower than the actual number.

The characteristics of vehicles are also a source of mea-

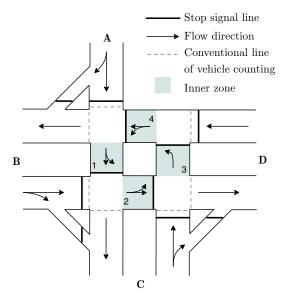


Fig. 1. The experimental junction

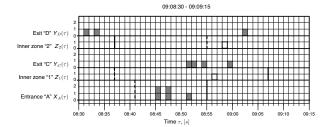


Fig. 2. Graphical representation of OD vehicle cows ACf and ADf.

surement errors. High vehicles passing in front of camera will hide the smaller vehicles or the whole camera Geld, i.e. they will produce a masking effect. Two-wheeled vehicles are only seldom counted because they are small. The heterogeneous colours of vehicle roofs also add to the problem of detection.

Meteorological conditions inevitably have an incuence on all types of trafte measurements and the video are blurred: the wind shakes the posts the cameras are Exed to, the sun s rays cause the recections on the vehicle surfaces and camera lenses, rain, snow and fog obscure a camera Geld. Changes in brightness caused by the position of the sun, clouds and headlights at night also determine the reliability of the measurements.

III. PROBLEM STATEMENT

Let $x_i(c)$ be the ∞ volume at entrance i during a trafter light cycle c and $y_i(c)$ be the ∞ volume which entered the junction during cycle c and leaves it by exit j. OD αow rate $b_{ij}(c)$ is the proportion of the ∞ of vehicles that come from entrance i and go to exit j. The problem is to estimate OD omega rates $b_{ij}(c)$ ($\forall i \in [1, n], \forall j \in [1, m]$) at the end of each traf light cycle c, such that

$$y_{j}(c) = \sum_{i=1}^{n} b_{ij}(c)x_{i}(c),$$

$$x_{i}(c)b_{ij}(c) \ge z_{kij}(c) \quad \forall k \text{ s.t. } \delta_{kij} = 1,$$

$$\sum_{j=1}^{m} b_{ij}(c) = 1,$$

$$(1a)$$

$$x_i(c)b_{ij}(c) \ge z_{kij}(c) \quad \forall k \text{ s.t. } \delta_{kij} = 1,$$
 (1b)

$$\sum_{i=1}^{m} b_{ij}(c) = 1,$$
 (1c)

$$b_{ij}(c) \ge 0, \tag{1d}$$

where $z_{kij}(c)$ is the number of vehicles which cross the junction from i to j and stop at inner zone k during cycle c, $\delta_{kij} = 1$ if OD ∞ from i to j can pass through inner zone k and is 0 otherwise. For a given cycle c the value of variable $x_i(c)$ can be obtained from instantaneous vehicle counts

$$x_i(c) = \sum_{\tau=1}^{\mathcal{G}_i(c)} \mathcal{X}_i(\tau),$$

where $\mathcal{G}_i(c)$ is a duration of the green light of cycle c in entrance i. The values of $y_i(c)$ and $z_{kij}(c)$ cannot be obtained directly from trafte measurements, because it is impossible to know the period of time during which the vehicle $\cos x_i(c)$ leaves the junction or stops at inner zones.

In order to obtain the values of $y_j(c)$ and $z_{kij}(c)$ from $\mathcal{Y}_j(\tau)$ and $\mathcal{Z}_k(\tau)$ respectively, we have proposed a dynamical model of vehicle cows crossing the junction using High-Level Petri Nets [13]. This model provides the onsets of the beginnings and the ends of the ∞ ws. For a given cycle c we thus know the possible duration of the presence of the cows in each zone and can collect the corresponding measurements during the same period. Therefore it is possible to put into a oneto-one correspondence the cows crossing the junction zones and the measurements taken at this zones. Consequently, at each cycle c, we can build dynamically a vehicle conservation law, represented by the following system of equations whose unknowns are the OD cow rates:

$$\mathbf{y} = \mathbf{X_1}\mathbf{b},\tag{2}$$

such that

$$X_2b \ge z,$$
 (3a)

$$\mathbf{Ib} = \mathbf{1},\tag{3b}$$

$$\mathbf{b} > 0, \tag{3c}$$

where y and z are vectors composed of J and K elements respectively, **b** is a vector containing the I unknowns b_{ij} , $\mathbf{X_1}$ is an $J \times I$ matrix rearranged so that (2) is equal to (1a), $\mathbf{X_2}$ is an $K \times I$ matrix built so that (3a) is equivalent to (1b), **I** is an indicator $n \times I$ matrix organised in such a way that (3b) is equal to (1c) and $\mathbf{1}$ is an identity n-vector. Note that for the clarity of notation the cycle c has been omitted.

The number of unknowns I depends on the junction structure and on the number of platoons that compose an OD cow. Depending on the trafte control strategy and the cow volume, an OD oow can leave the junction as a single platoon or as several platoons. The number of equations J depends on the number of exits and, in the same manner, on the control strategy. Since $I \geq J$ (in this paper $I \approx 10$), the system of equations (2) is underdetermined.

In order to overcome the problem of the system indetermination and to guarantee the uniqueness of the solution it is necessary to use some additional information about the values of the OD cow rates. In this paper, for choosing a solution among a solution set, we propose to use the mean value of the OD ∞ rates estimated during a time period preceding the estimation, called *sliding mean value*. Moreover, to model the uncertainty and imprecision of the real data, we suggest to introduce slack variables and to represent the values of b, X_1, X_2, y and z, Est, by intervals, and, second, by triangular fuzzy numbers.

In Cine, four methods are proposed in this paper. To estimate the crisp unknowns b when the elements of model (2) and of constraint (3c) are crisp, we apply the ordinary least squares (LS) and the linear programming (CLP) methods using the sliding mean value of b. Linear programming (*ILP*) and fuzzy linear programming (FLP) approaches are formulated for the cases where these elements are represented by intervals and fuzzy numbers respectively. We introduce in the sequel the four methods and show their results on our application.

IV. OD FLOW RATES ESTIMATION

A. Least squares method with crisp variables and coefŒients (CLS)

The estimation problem of OD ∞ wrates b can be considered as an estimation of coefGeients of the regression model equivalent to the relationship (2). Let the input and output variables $\mathbf{X_1}, \mathbf{X_2}, \mathbf{y}, \mathbf{z}$ and the coefGeients b be non-negative crisp numbers. A Gest approach to estimate b is to use a least squares method. A sliding mean values $\overline{\mathbf{b}} = \{\overline{b}_l\}_{l=1,\dots,8}$, available for eight OD pairs which are described in Section II, are used to provide a unique solution. To insure the existence of feasible solutions, we introduce slack variables, $e_{1,j}$ ($\forall j=1,\dots,J$) and $e_{1,k}$ ($\forall k=1,\dots,K$), and we propose to solve the following problem:

$$\min_{\mathbf{b}^*, \mathbf{b}, \xi, \mu, \eta} \left(\| \mathbf{b}^* - \overline{\mathbf{b}} \|^2 + \gamma_1 \sum_{j=1}^{J} e_{1,j} + \gamma_2 \sum_{j=1}^{J} e_{2,j} + \gamma_3 \sum_{k=1}^{K} e_{3,k} \right)$$
(4)

s.t.

$$\begin{cases}
\mathbf{Ab} = \mathbf{b}^*, \\
\mathbf{Ib}^* = \mathbf{1}, \\
\mathbf{X_1b} = \mathbf{y} + \mathbf{e_1} - \mathbf{e_2}, \\
\mathbf{X_2b} + \mathbf{e_3} \ge \mathbf{z}, \\
\mathbf{b} \ge 0, \\
\mathbf{e_1} \ge 0, \mathbf{e_2} \ge 0, \mathbf{e_3} \ge 0,
\end{cases}$$
(5)

where γ_1 , γ_2 and γ_3 are given weight coefteients, \mathbf{b}^* is the vector of the eight OD \mathbf{x} was rates, considered in Section II, and \mathbf{A} is a matrix rearranged in such a way that $\mathbf{A}\mathbf{b} = \mathbf{b}^*$.

B. Linear programming problem with crisp variables and coefŒients (CLP)

A linear programming approach can be applied to estimate the crisp elements of the regression model (2). Using the same notations, we can build the following minimisation problem:

$$\min_{\mathbf{b}^*, \mathbf{b}, \mathbf{e_1}, \mathbf{e_2}, \mathbf{e_3}, \mathbf{e_4}, \mathbf{e_5}} \left(\gamma_1 \sum_{j=1}^{J} e_{1,j} + \gamma_2 \sum_{j=1}^{J} e_{2,j} + \gamma_3 \sum_{k=1}^{K} e_{3,k} + \gamma_4 \sum_{i=1}^{I} e_{4,i} + \gamma_5 \sum_{i=1}^{I} e_{5,i} \right)$$
(6)

s.t.

$$\begin{cases}
X_{1}b = y + e_{1} - e_{2}, \\
X_{2}b + e_{3} \ge z, \\
Ab = b^{*}, \\
b^{*} = \overline{b} + e_{4} - e_{5}, \\
Ib = 1, \\
b \ge 0, \\
e_{1} > 0, e_{2} > 0, e_{3} > 0, e_{4} > 0, e_{5} > 0.
\end{cases}$$
(7)

where $\gamma_1, \gamma_2, \gamma_3, \gamma_4$ and γ_5 are given weight coeftents and $e_{4,i}, e_{5,i} \ (\forall i=1,\ldots,I)$ are slack variables.

C. Linear programming problem with variables and coefŒ cients represented as intervals (ILP)

In order to take into account the inherent imprecision of the data, we choose to represent the elements of the model (2) and constraint (3a) by intervals, denoted as $a = [a^-, a^+]$. The values of X_1^- and X_1^+ (X_2^- and X_2^+) was inferred from empirical distributions of the error counts (cf Fig. 3). In most cases, counts of vehicles at the entrances are smaller than the true value of X_1 (X_2). Here, the equation (2) is interpreted as $X_1b \subseteq y$ and inequality constraint (3a) as the least conservative inequality $\mathbf{X}_{\mathbf{2}}^{+}\mathbf{b}^{+} \geq \mathbf{z}^{-}$. Since there is no available histograms of error counts for y and z, intervals $[y^-, y^+]$ and $[\mathbf{z}^-, \mathbf{z}^+]$ was constructed by adding to the known values of y and z the symmetrical spreads chosen experimentally. With respect to constraints (2)-(3c), we propose to estimate b such that $\mathbf{b}^- \leq \mathbf{b} \leq \mathbf{b}^+$ and the width of intervals $[\mathbf{b}^-, \mathbf{b}^+]$ tend toward zero. Considering that all variables and coefcients of the regression model are non-negative, the minimisation problem is written as follows:

$$\min_{\mathbf{b}, \mathbf{b}^{+}, \mathbf{b}^{-}, \mathbf{e_{1}}, \mathbf{e_{2}}, \mathbf{e_{3}}} \left(\sum_{i=1}^{I} \left(\mathbf{b}_{i}^{+} - \mathbf{b}_{i}^{-} \right) + \gamma_{1} \sum_{j=1}^{J} e_{1,j} + \gamma_{2} \sum_{j=1}^{J} e_{2,j} + \gamma_{3} \sum_{k=1}^{K} e_{3,k} \right)$$

$$(8)$$

s.t.

$$\begin{cases}
\mathbf{X_{1}^{-}b^{-}} \leq \mathbf{y}^{-} + \mathbf{e_{1}}, \\
\mathbf{X_{1}^{+}b^{+}} + \mathbf{e_{2}} \geq \mathbf{y}^{+}, \\
\mathbf{X_{2}^{+}b^{+}} + \mathbf{e_{3}} \geq \mathbf{z}^{-}, \\
\mathbf{Ib} = \mathbf{1}, \\
0 \leq \mathbf{b}^{-} \leq \mathbf{b} \leq \mathbf{b}^{+} \leq 1, \\
\mathbf{e_{1}} \geq 0, \mathbf{e_{2}} \geq 0, \mathbf{e_{3}} \geq 0,
\end{cases}$$
(9)

where γ_1 , γ_2 and γ_3 are given weight coef@cients.

D. Linear programming problem with fuzyy variables and coefŒients (FLP)

Let $\tilde{\mathbf{X}}_1, \tilde{\mathbf{X}}_2, \tilde{\mathbf{y}}, \tilde{\mathbf{z}}$ and $\tilde{\mathbf{b}}$ be triangular fuzzy numbers represented as triple $\tilde{\mathbf{a}} = (\mathbf{a}^-, \mathbf{a}^{\mathbf{m}}, \mathbf{a}^+)$. The form of fuzzy numbers $\tilde{\mathbf{X}}_1$ was derived from empirical distributions of the error counts as shown in Fig. 3. The fuzzy number $\tilde{\mathbf{X}}_2$ is determined similarly to $\tilde{\mathbf{X}}_1$. Because of the lack of histograms of error counts for \mathbf{y} and \mathbf{z} , fuzzy numbers $\tilde{\mathbf{y}}$ and $\tilde{\mathbf{z}}$ are supposed to be symmetrical with the spreads chosen experimentally.

One of the way for interpreting the equality of two fuzzy numbers \tilde{a} and \tilde{b} (cf [15]) is to consider ${}^{\alpha}A\subseteq {}^{\alpha}B$ ($\forall \alpha\in[0,1]$), where the notation ${}^{\alpha}A=\{x|\mu_A(x)\geq\alpha\}$ stands for the α -cut of fuzzy number $\tilde{\bf a}$. In the same manner we decrease $\tilde{a}\geq\tilde{b}$ as ${}^{\alpha}A\geq{}^{\alpha}B$, $\forall \alpha\in[0,1]$.

According to (2)-(3c) and supposing the unknown OD ∞ rates $\dot{\mathbf{b}}$ to be triangular fuzzy numbers of the form $(\mathbf{b}^-, \mathbf{b}^m, \mathbf{b}^+)$, we propose to solve the previous problem (*ILP*) for each α -cut. For this reason a new constraint is introduced:

$$[\alpha_i \mathbf{b}^-, \alpha_i \mathbf{b}^+] \subset [\alpha_j \mathbf{b}^-, \alpha_j \mathbf{b}^+],$$

where i < j, $\forall i, j \in [1, N]$, and N is the Unite number of α -cuts. The following fuzzy linear programming is written:

$$\min_{\mathbf{b}^{m}, \mathbf{b}^{+}, \mathbf{b}^{-}, \mathbf{e}_{1}, \mathbf{e}_{2}, \mathbf{e}_{3}} \left(\sum_{\alpha=1}^{N} \sum_{i=1}^{I} \left({}^{\alpha} \mathbf{b}_{i}^{+} - {}^{\alpha} \mathbf{b}_{i}^{-} \right) + \gamma_{1} \sum_{j} e_{1,j} + \gamma_{2} \sum_{j} e_{2,j} + \gamma_{3} \sum_{k} e_{3,k} \right) (10)$$

s.t.

t.
$$\begin{cases}
\alpha \mathbf{X}_{1}^{-\alpha} \mathbf{b}^{-} \leq \alpha \mathbf{y}^{-} + \mathbf{e}_{1}, \\
\alpha \mathbf{X}_{1}^{+\alpha} \mathbf{b}^{+} + \mathbf{e}_{2} \geq \alpha \mathbf{y}^{+}, \\
\alpha \mathbf{X}_{2}^{+\alpha} \mathbf{b}^{+} + \mathbf{e}_{3} \geq \alpha \mathbf{z}^{-}, \\
0 \leq \alpha_{i} \mathbf{b}^{-} \leq \mathbf{b}^{m} \leq \alpha_{i} \mathbf{b}^{+} \leq 1, \quad \forall i \in [1, N] \\
\alpha_{i} \mathbf{b}^{-} \leq \alpha_{j} \mathbf{b}^{-} \leq \alpha_{j} \mathbf{b}^{+} \leq \alpha_{i} \mathbf{b}^{+}, \quad i < j, \forall i, j \in [1, N], \\
\mathbf{I} \mathbf{b}^{m} = \mathbf{1}, \\
\mathbf{e}_{1} \geq 0, \mathbf{e}_{2} \geq 0, \mathbf{e}_{3} \geq 0,
\end{cases}$$
(11)

where γ_1 , γ_2 and γ_3 are given weight coef $\mathbf{\tilde{x}}$ ents and the α -cuts of the fuzzy numbers $\mathbf{\tilde{X}_1}$, $\mathbf{\tilde{X}_2}$, $\mathbf{\tilde{y}}$ and $\mathbf{\tilde{z}}$ are calculated as follows:

$$\begin{array}{l} ^{\alpha}\mathbf{X}_{1}^{-}=\alpha(\mathbf{X}_{1}^{m}-\mathbf{X}_{1}^{-})+\mathbf{X}_{1}^{-}, \quad ^{\alpha}\mathbf{X}_{1}^{+}=\alpha(\mathbf{X}_{1}^{m}-\mathbf{X}_{1}^{+})+\mathbf{X}_{1}^{+}, \\ ^{\alpha}\mathbf{X}_{2}^{+}=\alpha(\mathbf{X}_{2}^{m}-\mathbf{X}_{2}^{+})+\mathbf{X}_{2}^{+}, \quad ^{\alpha}\mathbf{y}^{-}=\alpha(\mathbf{y}^{m}-\mathbf{y}^{-})+\mathbf{y}^{-}, \\ ^{\alpha}\mathbf{y}^{+}=\alpha(\mathbf{y}^{m}-\mathbf{y}^{+})+\mathbf{y}^{+}, \quad ^{\alpha}\mathbf{z}^{-}=\alpha(\mathbf{z}^{m}-\mathbf{z}^{-})+\mathbf{z}^{-}. \end{array}$$

V. EXPERIMENTAL RESULTS

These methods have been tested using real data collected at the experimental junction during 30-minutes period with peak trafte conditions. The estimation is made on 25 consecutive trafte light cycles. The actual values of OD ∞ rates β_i ($\forall i=1,\ldots,8$), calculated manually from video images, are available for all cycles.

The width of intervals \mathbf{y} and \mathbf{z} of the *ILP* method and the support width of fuzzy numbers $\tilde{\mathbf{y}}$ and $\tilde{\mathbf{z}}$ of the *FLP* method were experimentally Exed to 2. The weight coef Exients in all models has been Exed as follows: $\gamma_1 = 0.1$, $\gamma_2 = 0.1$, $\gamma_3 = 0.2$, $\gamma_4 = \gamma_5 = 1$.

The estimation error has been calculated for the OD ∞ w rates (Fig. 4): $E = \hat{\mathbf{b}}\mathbf{X}^* - \beta \mathbf{X}^*$, where \mathbf{X}^* is a vector of actual vehicle counts at the entrances of the experimental junction, $\hat{\mathbf{b}}$ is a vector of estimated OD ∞ w rates. For *FLP* method we

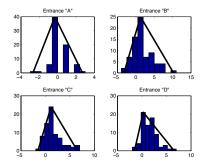


Fig. 3. Empirical distributions of error counts for the entrances of the junction

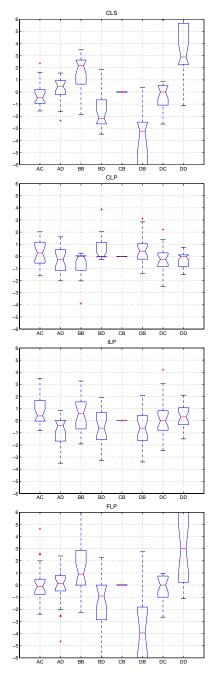


Fig. 4. Estimation errors (in number of vehicles) for 8 OD ∞ w rates for the 25 cycles

consider $\hat{\mathbf{b}} = \mathbf{b}^m$. The best results are obtained with *CLP* and *FLP* methods for which the median error is about zero for all OD ∞ ws. The prediction of the exit ∞ ws with *ILP* method is less accurate than those of the *FLP* method (cf Fig. 5). Note that the estimation error E of all methods is higher, if the ∞ w volume is lower, like for the OD ∞ ws ADf and BBf.

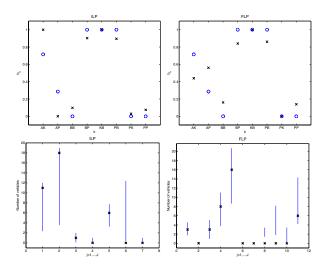


Fig. 5. Estimations $\hat{\mathbf{b}}$ (x) and true values β (circles) for 8 OD ∞ rates and model prediction $[\mathbf{X}_1^-\hat{\mathbf{b}},\mathbf{X}_1^+\hat{\mathbf{b}}]$ of exit ∞ \mathbf{y} for one particular cycle

VI. CONCLUSION

A new short-time estimation method of the OD matrix for a signalised junction have been proposed in this paper. The method is founded on the construction of a conservation law of vehicles at each trafter light cycle, represented by the underdetermined system of equations. This system is obtained in a dynamical way from the model of trafter cows built using the High-Level Petri Nets (HLPN) [13]. Real data collected at the experimental signalised junction the detection of the control of the control

Four approaches have been proposed to estimate the OD matrix. The inherent imprecision of the data was modelled in three different ways: by introducing slack variables, by representing the data by intervals and by triangular fuzzy numbers. An additional historical knowledge about the values of the OD ∞ rates is used in two of the proposed approaches to guarantee the uniqueness of the solution.

The best results were obtained with two linear programming approaches, one of which is based on the use of the sliding mean value of the OD ∞ rates and the second is founded on fuzzy modelling of the data. The latest method is applied if the lack of precision of trafter measurements is important. However, this method does not guarantee the uniqueness of the solution.

Our future lines of research will be centred, Œst, on the improvement of the FLP method in order to provide the best and unique estimation of the OD cow rates and, second, on the modelling of the temporal uncertainty of the data by introducing the notion of a fuzzy token in the HLPN model. The application of our method should be also extended to a sequence of junctions.

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